Sequence analysis

Partial correlation coefficient between distance matrices as a new indicator of protein-protein interactions

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ABSTRACT

Motivation: The computational prediction of protein-protein interactions is currently a major issue in bioinformatics. Recently, a variety of co-evolution-based methods have been investigated toward this goal. In this study, we introduced a partial correlation coefficient as a new measure for the degree of co-evolution between proteins, and proposed its use to predict protein-protein interactions.

Results: The accuracy of the prediction by the proposed method was compared with those of the original mirror tree method and the projection method previously developed by our group. We found that the partial correlation coefficient effectively reduces the number of false positives, as compared with other methods, although the number of false negatives increased in the prediction by the partial correlation coefficient.

Availability: The R script for the prediction of protein-protein interactions reported in this manuscript is available at http://timpani.genome. ad.jp/~parco/

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Supplementary information: The information is also available at the same site as the R script.

1 INTRODUCTION

The use of co-evolutionary information is one of the popular approaches for predicting a protein-protein interaction (PPI). This approach is based on the assumption that interacting protein pairs are likely to evolve in a correlated fashion, and that the phylogenetic trees of the interacting proteins are similar to each other, owing to the co-evolution. Based on several previous reports, such as that by Goh et al. (2000), Pazos and Valencia (2001) developed a method to evaluate the intensity of co-evolution by comparing a pair of distance matrices, or a set of genetic distances, instead of phylogenetic trees. The intensity of co-evolution is calculated with the Pearson's correlation coefficient between the distance matrices (for details, see Methods). This is referred to as the 'mirror tree' method, according to the assumption. In recent years, some variants of the mirror tree method have been proposed (Gertz et al., 2003; Goh and Cohen, 2002; Kim et al., 2004; Ramani and Marcotte, 2003; Tan et al., 2004). One of the latest improvements in the mirror tree framework is the reduction of false positives, which is facilitated by removing the information about the evolutionary relationships of the organism sources from the distance matrices. It has been pointed out that predictions by the mirror tree methods tend to introduce many false positives. One explanation for this phenomenon is that, in the mirror tree method, the distance matrices of all of the proteins are constructed with a set of orthologues from the same organism sources. As a result, all of the distance matrices share the information about the evolutionary relationship of the sources, which is suspected of being the cause of the false positives. Sato et al. (2005) and Pazos et al. (2005) independently developed methods to remove such evolutionary information from the distance matrices, and demonstrated that the approaches can dramatically reduce the number of false positives in the prediction.

In this manuscript, we propose the use of the partial correlation coefficient as a novel measure for the intensity of co-evolution, to improve the original mirror tree approach. From the definition (see Methods), the partial correlation coefficient is expected to be useful to remove the information about the evolutionary relationship of the sources. The performance of the partial correlation coefficient, in terms of reducing the false positives from the prediction, was improved, as compared with that reported by Sato et al. (2005). We also examined the relationship between the accuracy of the PPI prediction by the co-evolutionary analysis and the number of proteins used in the study.

2 METHODS

2.1 Dataset

We selected 821 interacting pairs of Escherichia coli proteins that were experimentally determined, from the Database of Interacting Proteins (DIP) Version 01/16/2006 (Salwinski et al., 2004). The selected pairs did not include homo-interactions, and comprised 179 different E.coli proteins. All of the proteins are registered in KEGG/KO (Kanehisa et al., 2006), a database of putative orthologous sequences identified from fully sequenced genome data. Then, putative orthologues corresponding to 179 proteins derived from E.coli were collected from 65 different bacterial species, according to the description in KEGG/KO. Hereafter, the putative orthologues are simply referred to as the orthologues. The lists of PPI pairs, protein names and source organisms are provided in the Supplementary

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data. One of the assumptions in this study is that a pair of proteins from a bacterium, which are each orthologous to the interacting proteins of *E.coli*, also interacts in the bacterial species. The other assumption is that the interaction affects the co-evolution of the orthologues.

2.2 Multiple sequence alignment and distance matrix

A multiple alignment of each set of selected orthologous proteins was performed with the alignment software, MAFFT (Katoh *et al.*, 2005). A distance matrix for the orthologues was calculated from the multiple sequence alignment. Suppose that we want to obtain a distance matrix, D_A , for protein A. The orthologues of protein A are collected from *n* species. The size of D_A is $n \times n$, and each row or column of the matrix corresponds to a species under consideration. An element of the matrix, $D_A(x, y)$, represents the genetic distance of protein A between species *x* and *y*. The distance was calculated as a maximum likelihood estimate, with the PROTDIST module in the PHYLIP package (Felsenstein, 2004). The score table by Jones *et al.* (1992) was used for the maximum likelihood estimation. The distance matrix is symmetric, and only the upper or lower triangular part of the matrix is used in the prediction process.

2.3 Transformation from distance matrix to phylogenetic vector

The distance matrix was transformed into a vector for easier formulation. The upper or lower half of the non-diagonal elements of the distance matrix was arranged as an array of the numerical values in a certain order. This operation was applied to all of the proteins, and the corresponding distance matrices were transformed into vectors with the same order of the elements. When the matrix has a size of $n \times n$, the dimension of the vector is n(n-1)/2. The vector is hereafter referred to as a 'phylogenetic vector'. As described below, n was set to 66 in this study. Let us consider the pair of protein i and protein j, and their phylogenetic vectors $|v_i\rangle$ and $|v_j\rangle$, transformed from the corresponding distance matrices D_i and D_j , where the subscripts i and j indicate different sets of orthologues. Then, we applied the normalization of the elements of each vector with the average and the standard deviation of the elements as follows:

$$|\nu_i^{\star}\rangle = \frac{|\nu_i\rangle - |\mu\rangle}{\sqrt{\operatorname{Var}(\nu_i)}},\tag{1}$$

where $|\mu\rangle$ is a vector with the same dimension as $|\nu_i\rangle$. All of the elements of $|\mu\rangle$ are constant, and are equivalent to the arithmetic average over the elements of $|\nu_i\rangle$. Var (ν_i) indicates the variance over all of the elements of $|\nu_i\rangle$. The inner product of a pair of normalized vectors is reduced to the Pearson's correlation coefficient used for the mirror tree method. Hereafter, the correlation coefficient is denoted as ρ_{ii}^{MIRROR} .

$$\rho_{ii}^{\text{MIRROR}} = \langle \nu_i^{\star} \mid \nu_i^{\star} \rangle. \tag{2}$$

2.4 Projection operator

Here we will briefly review the projection method proposed in our previous manuscript (Sato *et al.*, 2005). In this method, we use a unit vector $|u\rangle$, which represents the evolutionary relationship of the species. We designed the unit vector in three different ways: (1) transforming the distance matrix of 16S ribosomal RNA (rRNA) from the same source organisms as for the proteins under consideration, (2) averaging the phylogenetic vectors and (3) extracting the principal components of the phylogenetic vectors. The unit vectors obtained in these three ways were respectively designated as $|u_{16S}\rangle$, $|u_{AVE}\rangle$ and $|u_{PC1}\rangle$.

Using the unit vector $|u\rangle$, a projection operator P is obtained as

$$P = I - |u\rangle\langle u|,\tag{3}$$

where $|u\rangle\langle u|$ is also a projection operator onto the direction of the unit vector $|u\rangle$. *I* represents an identity matrix with the same size as $|u\rangle\langle u|$. By applying

the projection operator (3) to a phylogenetic vector, say, $|\nu_i\rangle$, the component orthogonal to $|\mu\rangle$ is obtained as follows:

$$|\varepsilon_i\rangle = P|\nu_i\rangle = |\nu_i\rangle - |u\rangle\langle u|\nu_i\rangle. \tag{4}$$

The projection operator can exclude the information about the evolutionary relationship among the source organisms from a phylogenetic vector. The same projection operator was applied to all of the phylogenetic vectors under consideration. Each of the residual vectors defined by formula (4) was normalized with the average and the standard deviation of the elements. Consider a pair of normalized vectors, $|\varepsilon_i^*\rangle$ and $|\varepsilon_j^*\rangle$. Then, the inner product of the two vectors

$$\rho_{ii}^{\text{PROJECTION}} = \langle \boldsymbol{\varepsilon}_i^{\star} | \boldsymbol{\varepsilon}_i^{\star} \rangle \tag{5}$$

represents the Pearson's correlation coefficient between the residues, after excluding the information about the evolutionary relationship from the original phylogenetic vectors. The Pearson's correlation coefficients between the residual vectors for proteins *i* and *j*, based on the unit vectors $|u_{16S}\rangle$, $|u_{AVE}\rangle$ and $|u_{PC1}\rangle$ in the construction of the projection operator *P*, are represented by ρ_{ij}^{IGS} , ρ_{ij}^{AVE} and ρ_{ij}^{PC1} , respectively.

2.5 Partial correlation coefficient

Suppose that we have *m* proteins and we want to detect interacting pairs among them. Let us consider multiple regressions of $|\nu_i\rangle$ and $|\nu_j\rangle$ with (m-2) phylogenetic vectors as follows:

$$|\nu_i\rangle = \alpha_0 + \sum_{k \neq i,j}^m \alpha_k |\nu_k\rangle + |\delta_i\rangle, \tag{6}$$

$$|\nu_{j}\rangle = \beta_{0} + \sum_{l\neq i,j}^{m} \beta_{l} |\nu_{l}\rangle + |\delta_{j}\rangle, \qquad (7)$$

where α_i and β_i are scalar parameters. The linear combination of (m-2)phylogenetic vectors is expected to represent the evolutionary relationship of the source organisms, since the principal component analysis of a number of phylogenetic vectors suggested that the first principal component vector, with a cumulative rate of contribution >80%, represents the evolutionary relationship (data not shown). Therefore, the residual vectors, $|\delta_i\rangle$ and $|\delta_i\rangle$, essentially lack the evolutionary information of the source organisms. $|v_i\rangle$ is excluded from the summation on the right side of the first equation. If protein *i*, represented by $|\nu_i\rangle$, co-evolves with protein *j*, represented by $|\nu_i\rangle$, through a PPI, then the effect of the co-evolution with protein j is expected to be present in the residual vector $|\delta_i\rangle$. Likewise, the effect of co-evolution with protein *i*, if any, is expected to be present in $|\delta_i\rangle$. Therefore, the similarity between the two residual vectors is considered to indicate the extent of co-evolution between proteins *i* and *j*. To evaluate the similarity between the residual vectors, we normalized the vectors with the averages and the standard deviations, and we expressed them as $|\delta_i^{\star}\rangle$ and $|\delta_i^{\star}\rangle$. The inner product between the normalized residual vectors is a measure to indicate the similarity between them, and is called the partial correlation coefficient. That is, the partial correlation coefficient $\rho_{ij}^{\text{PARTIAL}}$ between $|\nu_i\rangle$ and $|\nu_j\rangle$ is expressed as follows:

$$\rho_{ij}^{\text{PARTIAL}} = \langle \delta_i^{\star} | \delta_j^{\star} \rangle. \tag{8}$$

Instead of constructing the above multiple regression models in each computation, the following formula was used to obtain the partial correlation coefficient.

$$\rho_{ij}^{\text{PARTIAL}} = \frac{-(R^{-1})_{ij}}{\sqrt{(R^{-1})_{ii}(R^{-1})_{jj}}},\tag{9}$$

where *R* is the correlation coefficient matrix whose (i, j)-th element is ρ_{ij}^{MIROR} , and the superscript -1 indicates the inverse of the matrix. For the derivation of Equation (9), see Supplementary data. When the subscripts, *i* and *j*, are omitted, $\rho^{\#}$ collectively represents the type of correlation coefficient indicated by the superscript.

Table 1. Comparison of the top 20 protein pairs sorted in decreasing order of the correlation coefficients

| Rank | ρ^{MIRROR} | | $ ho^{16\mathrm{S}}$ | | $ ho^{ m AVE}$ | | $ ho^{ m PC1}$ | | $ ho^{\mathrm{PARTIAL}}$ | |
|------|------------------------|--------------|----------------------|--------------|----------------|--------------|----------------|--------------|--------------------------|--------------|
| 1 | rpoC - rpoB | 0.9883* | sdhB - sdhA | 0.9659* | sdhB - sdhA | 0.9562* | sdhB - sdhA | 0.9586* | nrdB – nrdA | 0.9360* |
| 2 | rpoC - dnaN | 0.9831 | priA - ileS | 0.9423 | priA - ileS | 0.9291 | nrdB - nrdA | 0.9268^{*} | sdhB – sdhA | 0.9103* |
| 3 | sdhB - sdhA | 0.9820^{*} | carB - murC | 0.9398 | nrdB - nrdA | 0.9271* | priA - ileS | 0.9242 | trpB - trpA | 0.6719* |
| 4 | rpoC - polA | 0.9814^{*} | rpoC - rpoB | 0.9288^{*} | rpoD - rpoC | 0.9257^{*} | carB - murC | 0.9158 | gltD - gltB | 0.6708^{*} |
| 5 | groL - rpoC | 0.9794 | nrdB - nrdA | 0.9207^{*} | carB - murC | 0.9187 | rpoD - rpoC | 0.8953* | pstS - pstB | 0.5679 |
| 6 | priA - ileS | 0.9790 | rpoD - rpoC | 0.9103* | rpoC - rpoB | 0.8860^{*} | murC - trxB | 0.8685 | atpE - atpB | 0.5531* |
| 7 | rpoB - nusA | 0.9786^{*} | murC - trxB | 0.9075 | murC - trxB | 0.8737 | carB - trxB | 0.8542 | valS - gltA | 0.5263 |
| 8 | rpoB - rpoA | 0.9783* | rpoC - dnaN | 0.8976 | malG - malF | 0.8722^{*} | rpoC - rpoB | 0.8464* | groL - dnaK | 0.5217* |
| 9 | groL - polA | 0.9780 | carB - trxB | 0.8961 | carB - trxB | 0.8605 | malG - malF | 0.8386^{*} | hemC - murA | 0.5106 |
| 10 | carB - murC | 0.9777 | hflB - rho | 0.8925 | rpoD - gyrB | 0.8506 | rpoD - gyrB | 0.8375 | gltA - pyrD | 0.4598 |
| 11 | uvrC - uvrB | 0.9771^{*} | rpoC - nusG | 0.8889^{*} | priA - serS | 0.8438 | priA - serS | 0.8375 | malE - malG | 0.4474^{*} |
| 12 | dnaX - dnaG | 0.9765 | rpoB - nusA | 0.8752^{*} | rpoD - rpoB | 0.8408^{*} | serS - ileS | 0.8330 | sucC - sucD | 0.4373^{*} |
| 13 | rpoB - dnaN | 0.9763 | rpoD - rpoB | 0.8716^{*} | serS - ileS | 0.8398 | ileS - clpP | 0.8310 | rplK - valS | 0.4215 |
| 14 | rpoC - nusG | 0.9760^{*} | priA - serS | 0.8711 | rpoB - nusA | 0.8297^{*} | gltD - gltB | 0.8245^{*} | valS - leuS | 0.4039 |
| 15 | uvrA - dnaX | 0.9759 | serS - ileS | 0.8689 | ileS - clpP | 0.8293 | ileS - sucB | 0.8155 | leuS - purB | 0.3769 |
| 16 | rpoC - gyrB | 0.9756 | malG - malF | 0.8678^{*} | rpoC - dnaN | 0.8230 | sucC - sucD | 0.8043^{*} | hflB - rho | 0.3603 |
| 17 | hflB - rho | 0.9754 | prlA - polA | 0.8626 | ileS - sucB | 0.8229 | hflB - rho | 0.8030 | dnaK - serS | 0.3601 |
| 18 | rpoB - polA | 0.9746^{*} | ileS - clpP | 0.8555 | rpoD - dnaN | 0.8187 | rpoB - nusA | 0.7999^{*} | gltX - leuA | 0.3578 |
| 19 | rpoA - polA | 0.9745^{*} | rpoB - rpoA | 0.8550^{*} | gltD - gltB | 0.8157^{*} | priA - clpP | 0.7838 | infA - murA | 0.3441 |
| 20 | uvrA - lepA | 0.9745 | murC - glmS | 0.8548 | hflB - rho | 0.8141 | greA - pnp | 0.7838 | rpsN - ilvB | 0.3418* |

3 RESULTS AND DISCUSSION

3.1 Prediction of protein–protein interaction by co-evolutionary analysis

To evaluate the intensity of the co-evolution between proteins, we calculated five types of correlation coefficients, ρ^{MIRROR} , ρ^{168} , ρ^{AVE} , ρ^{PC1} and $\rho^{PARTIAL}$, for all of the possible pairs of 179 proteins, i.e. 15931 pairs of proteins. In this analysis, the unit vectors required for the calculations of ρ^{AVE} and ρ^{PC1} were obtained by the average operation or the principal component analysis of the phylogenetic vectors of the 179 proteins, whereas 177 proteins were used as the explanatory variables of multiple regression for the calculation of the partial correlation coefficient between the remaining two proteins. The correlation coefficients, sorted in decreasing order, are listed in the Supplementary data, and only the top 20 members of the lists are shown in Table 1. Interacting pairs registered in DIP are highlighted by asterisks in the table. As shown in the table, the top 20 positions for either method were occupied by 8 or 9 interacting protein pairs. In the case of ρ^{PARTIAL} , however, the interacting pairs were the most abundant near the top of the list, as compared to the other methods. The fifth position of the list for ρ^{PARTIAL} is occupied by the pair, pstS and pstB. There is no description about the interaction between these proteins in DIP. However, both proteins are involved in the phosphate transport system, according to their KEGG/KO annotations. Therefore, the proteins may interact with each other. Even if we use a high value, say 0.9, as a threshold for the correlation coefficient to predict a PPI, ρ^{MIRROR} will produce many pairs with high scores, including many non-interacting proteins, leading to more false positives in the actual prediction. This observation agrees with the results of our previous work (Sato et al., 2005). In contrast, the presence of interacting proteins in the top 20 list and the rapid decreases of ρ^{16S} , ρ^{AVE} , ρ^{PC1} and $\rho^{PARTIAL}$ guarantee the accuracy of prediction with the correlation coefficients, if the threshold is set at a sufficiently high value. Out of the

| Table 2. Sensitivity and Specificity of th | e prediction |
|--|--------------|
|--|--------------|

| Method | Sensiti | vity | | | Specificity | | | |
|-------------------------|---------|-------|-------|-------|-------------|--------|--------|--------|
| | 0.9 | 0.8 | 0.7 | 0.6 | 0.9 | 0.8 | 0.7 | 0.6 |
| ρ^{MIRROR} | 23.02 | 60.29 | 80.15 | 87.09 | 7.54 | 6.72 | 6.32 | 5.81 |
| ρ^{16S} | 0.49 | 2.68 | 7.31 | 13.28 | 57.14 | 22.68 | 13.57 | 8.94 |
| ρ^{AVE} | 0.37 | 1.34 | 3.29 | 6.58 | 60.00 | 45.83 | 27.27 | 18.31 |
| ρ^{PC1} | 0.24 | 0.85 | 2.31 | 5.24 | 50.00 | 41.18 | 26.03 | 18.38 |
| ρ^{PARTIAL} | 0.24 | 0.24 | 0.24 | 0.49 | 100.00 | 100.00 | 100.00 | 100.00 |

 $\begin{aligned} \text{Sensitivity} = (\text{True positive} / (\text{True positive} + \text{False negative})) \times 100\%, \text{Specificity} = (\text{True positive} / (\text{True positive} + \text{False positive})) \times 100\%. \end{aligned}$

four correlation coefficients, the decrease of ρ^{PARTIAL} was the steepest.

To examine the effect of the threshold values on the prediction accuracy, we evaluated the performance by introducing four different thresholds, 0.9, 0.8, 0.7 and 0.6, for the correlation coefficient in each method. The performances of the original mirror tree method and our four proposed methods were compared in terms of the sensitivity and the specificity, and the results are summarized in Table 2. In each case, when a pair of proteins had a correlation coefficient larger than a threshold, the pairs of proteins were predicted to interact with each other. Table 2 shows that ρ^{PARTIAL} provided extremely high specificity at any threshold, while ρ^{MIRROR} provided high sensitivity in all of the cases, except for the threshold = 0.9. The high specificities of ρ^{PARTIAL} mean that the use of a partial correlation coefficient helps to drastically reduce the false positives, as compared to those of ρ^{MIRROR} . This is an important feature for practical use in actual applications, because we would start the prediction with the part with high confidence; that is, from the top of the ranked list of predictions.

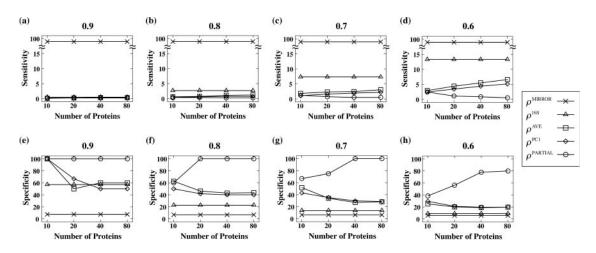


Fig. 1. The relationship between the prediction accuracy and the number of proteins. The x-axis indicates the number of proteins used in the analyses, and the y-axis indicates the sensitivity (a)-(d) or specificity (e)-(h). The performance of the five types of correlation coefficients was evaluated based on several threshold values, which are denoted at the top of each panel.

In other words, we would not be interested in the part of the prediction with lower confidence; that is, the lower ranking members of the prediction, in a practical situation. Table 2 also shows that the specificities of ρ^{16S} , ρ^{AVE} and ρ^{PC1} were much higher than those of ρ^{MIRROR} , but less than those of $\rho^{PARTIAL}$. On the other hand, the sensitivities of ρ^{16S} , ρ^{AVE} and ρ^{PC1} were higher than those of $\rho^{PARTIAL}$, but less than those of $\rho^{PARTIAL}$. These observations about ρ^{16S} , ρ^{AVE} and ρ^{PC1} agree with those in our previous report with regard to the sensitivity and the specificity of the correlation coefficients (Sato *et al.*, 2005). Thus, the partial correlation coefficient was more effective for the reduction of false positives from the prediction of PPI by the analysis of co-evolution than the previously proposed methods.

3.2 Prediction accuracy and the number of proteins

Out of the five types of correlation coefficients used in this study, it is clear from the definitions that ρ^{AVE} , ρ^{PC1} and $\rho^{PARTIAL}$ are dependent on the number of proteins used for the calculations. In contrast, ρ^{MIRROR} and ρ^{16S} are independent of the number of proteins. We examined the relationship between the prediction accuracy and the number of proteins by the following procedure.

- (1) Select *m* proteins randomly from the 179 proteins.
- (2) Compute the five types of correlation coefficients for every possible pair of the *m* proteins. As for ρ^{AVE} and ρ^{PC1} , the unit vector $|u\rangle$ was estimated with the phylogenetic vectors of the *m* proteins, whereas $\rho^{PARTIAL}$ for a pair was calculated with the remaining (m-2) proteins as explanatory variables.
- (3) Store the correlation coefficients in the five bins corresponding to the different types of correlation coefficients for each pair.
- (4) Repeat (1–3) until each bin for all of the possible pairs from the 179 proteins includes 100 correlation coefficients.

The mean value of 100 correlation coefficients in a bin was used as the corresponding type of correlation coefficient between a pair of proteins. When the mean correlation coefficient between two proteins was larger than a given threshold value, the proteins were predicted to interact with each other. We examined four cases of m, 10, 20, 30 and 40, in this study. For each case, the sensitivity and specificity for the five types of correlation coefficients were calculated. The sensitivity and the specificity are shown as the functions of the number of proteins in Figure 1.

The specificity for ρ^{PARTIAL} increases and those for ρ^{AVE} and ρ^{PC1} decrease, as the number of proteins increases. In contrast, ρ^{MIRROR} and ρ^{16S} have constant values, since they are independent from *m*, as described above. In any case, however, the specificity of ρ^{PARTIAL} is basically larger than or equal to those of other types of correlation coefficients. The difference in the dependence on the number of proteins between ρ^{PARTIAL} and other methods suggests that the prediction with ρ^{PARTIAL} is more useful for the reduction of false positives, when a large number of proteins is available for the prediction. At the same time, ρ^{PARTIAL} is usually effective in reducing the false positives, even when the number of proteins is small.

The sensitivities for ρ^{16S} , ρ^{AVE} , ρ^{PC1} and $\rho^{PARTIAL}$ were quite low, as compared to that of ρ^{MIRROR} . The sensitivities for ρ^{MIRROR} and ρ^{16S} were constant. In contrast, the sensitivities for the remaining three types of correlation coefficients showed a dependence on the number of proteins. As the number of proteins decreases, the sensitivity for $\rho^{PARTIAL}$ increases, whereas those for ρ^{AVE} and ρ^{PC1} decrease. In any case, the sensitivity of $\rho^{PARTIAL}$ was quite small and lower than those of the other types of correlation coefficients. Although the analyses suggested that the prediction with $\rho^{PARTIAL}$ includes more false negatives than other methods, the high specificity for $\rho^{PARTIAL}$ is considered to compensate for such deficits.

4 CONCLUSION

In this report, we demonstrated the potential of the partial correlation coefficient as a new measure for the intensity of co-evolution between proteins. In the numerical experiment, we showed that the performance of the partial correlation coefficient is better than or comparable with that of the projection method developed by Sato *et al.* (2005) and the original mirror tree method. We also examined the relationship between the prediction accuracy and the number of proteins used for the prediction analysis. The specificity for the prediction with ρ^{PARTIAL} was the highest among the five types of correlation coefficients, when a large number of sequences was available. Even when the number of proteins was small, the specificity for ρ^{PARTIAL} was better than those by the other methods. However, the prediction with ρ^{PARTIAL} included many false negatives, like the prediction by the projection methods. This problem should be examined for further improvement of PPI prediction by co-evolutionary analysis.

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Conflict of Interest: none declared.

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