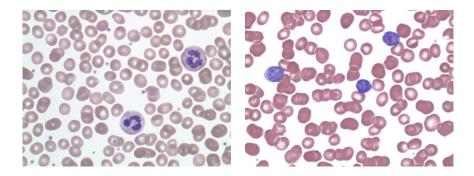
Fast sparse methods for genomics data

Jean-Philippe Vert



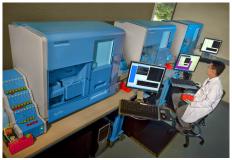
SMAI, Seignosse, May 28, 2013

### Normal vs cancer cells

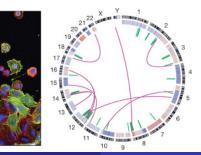


What goes wrong? How to treat?

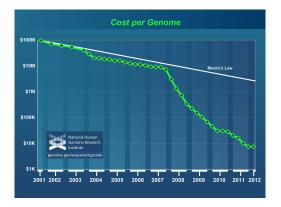
# Biology is now quantitative, "high-throughput"



DOE Joint Genome Institute



# Big data in biology



- "The \$1,000 genome, the \$1 million interpretation" (B. Kopf)
- High-dimensional, heterogeneous, structured data. "Large p"
- http://aws.amazon.com/1000genomes/

$$\min_{\boldsymbol{w}} \boldsymbol{R}(\boldsymbol{w}) + \lambda \Omega(\boldsymbol{w})$$

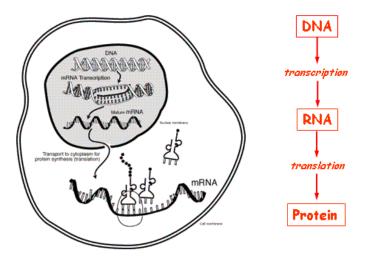
where:

- w is the hypothesis we want to infer from data
- *R*(*w*) is a smooth convex "fitness" function
- $\Omega(w)$  is a non-smooth convex penalty, which favors particular solution

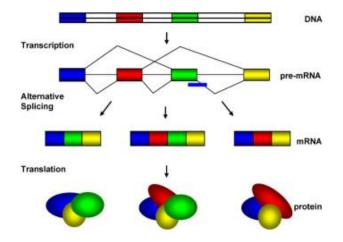
- Isoform detection from RNA-seq data
- 2 Mapping DNA breakpoints in cancer genomes



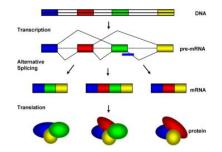
#### 2 Mapping DNA breakpoints in cancer genomes



# Alternative splicing: 1 gene = many proteins



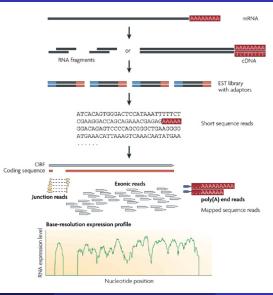
# The isoform identification and quantification problem



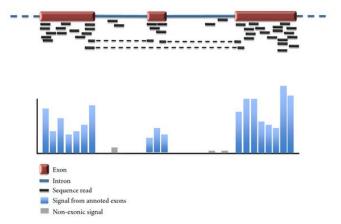
Given a biological sample (e.g., cancer tissue), can we:

- identify the isoform(s) of each gene present in the sample?
- Q quantify their abundance?

# RNA-seq measures mRNA abundance by sequencing short fragments

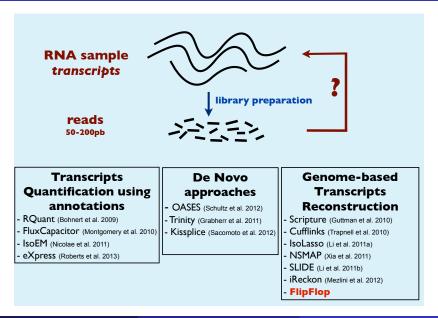


# RNA-seq and alternative splicing

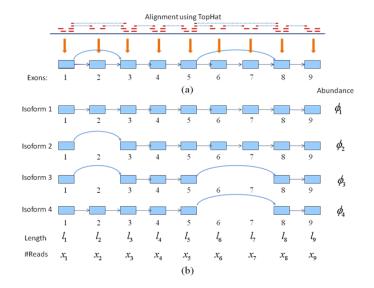


(Costa et al., 2011)

# From RNA-seq to isoforms



# The isoform deconvolution problem



(Xia et al., 2011)

*e* exons *c* candidate isoforms (up to  $2^e - 1$ )  $\phi \in \mathbb{R}^c_+$  the vector of abundance of isoforms (unknown!) *U* binary matrix:

	exon <sub>1</sub>		exon <sub>e</sub>	junction <sub>1,2</sub>	• • •	junction <sub>e1,e</sub>
isoform <sub>1</sub>	(1	•••	1	1	• • •	1 )
isoform <sub>2</sub>	1	•••	0	1	•••	0
:						
isoform <sub>c</sub>	0		1	0	• • •	o /

 $U^{\top}\phi$  the abundance of each exon/junction.

Goal: estimate  $\phi$  from the observed reads on each exon/junction

Estimate  $\phi$  sparse by solving:

$$\min_{\phi \in \mathbb{R}^{c}_{+}} \boldsymbol{R}(\boldsymbol{U}^{\top}\phi) + \lambda \| \phi \|_{1}$$

- IsoLasso (Li et al., 2011)
- NSMAP (Xia et al., 2011)
- SLIDE (Li et al., 2011)

Works well BUT computationally challenging to enumerate all candidate isoforms (up to 2<sup>*e*</sup>) for large genes!

# Fast isoform deconvolution with the Lasso

#### Theorem (Bernard, Mairal, Jacob and V., 2012)

The isoform deconvolution problem

$$\min_{\phi \in \mathbb{R}^{c}_{+}} R(U^{\top}\phi) + \lambda \|\phi\|_{1}$$

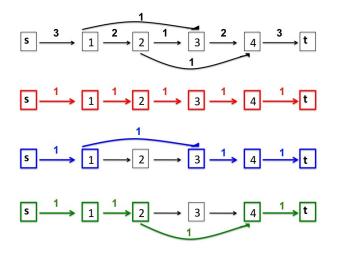
can be solved in polynomial time in the number of exon.

Key ideas

- $U^{\top}\phi$  corresponds to a flow on the graph
- Reformulation as a convex cost flow problem (Mairal and Yu, 2012)
- Recover isoforms by flow decomposition algorithm

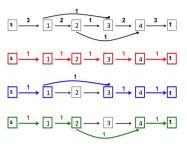
# "Feature selection on an exponential number of features in polynomial time"

### From isoforms to flows



- Isoforms are paths
- Linear combinations of isoforms are flows

# Isoform deconvolution as convex cost flow problem



$$\min_{\phi \in \mathbb{R}^{c}_{+}} R(U^{\top}\phi) + \lambda \| \phi \|_{1}$$

is equivalent to

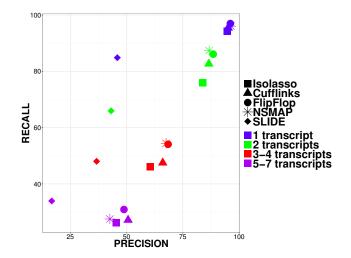
 $\min_{\text{f flow}} R(f) + \lambda f_t$ 

# FlipFlop (Bernard, Jacob, Mairal and V., 2012)

$$\min_{\phi \in \mathbb{R}^{c}_{+}} R(U^{\top}\phi) + \lambda \| \phi \|_{1}$$

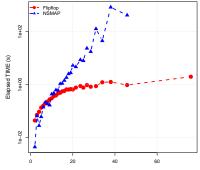
- Cufflink : a priori selection of isoforms (minimum graph cover)
- IsoLasso : pre-filtering of candidate isoforms using various heuristics
- NSMAP, SLIDE : limit the maximum number of exons
- FlipFlop : exact optimization without pre-filtering in polynomial time

# Performance in isoform identification

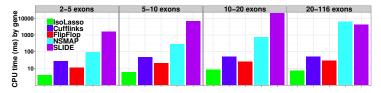


Simulated data (hg19, 1137 genes on chr1, 1million 75 bp single-end reads by transcript levels).

# Speed trial



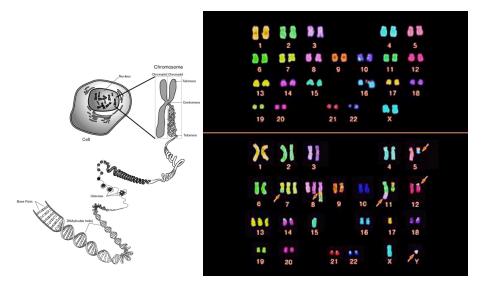
Number of EXONS







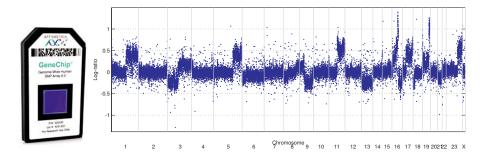
### Chromosomic aberrations in cancer



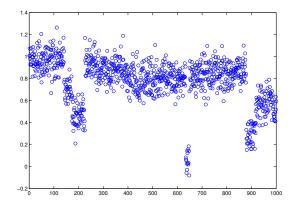
# Comparative Genomic Hybridization (CGH)

#### Motivation

- Comparative genomic hybridization (CGH) data measure the DNA copy number along the genome
- Very useful, in particular in cancer research to observe systematically variants in DNA content

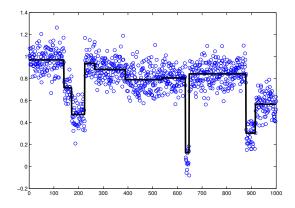


# Can we identify breakpoints and "smooth" each profile?

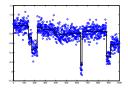


A classical multiple change-point detection problem
 Should scale to lengths of order  $10^6 \sim 10^9$ 

# Can we identify breakpoints and "smooth" each profile?



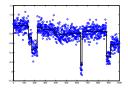
A classical multiple change-point detection problem
 Should scale to lengths of order  $10^6 \sim 10^9$ 



For a signal Y ∈ ℝ<sup>p</sup>, define an optimal approximation β ∈ ℝ<sup>p</sup> with k breakpoints as the solution of

$$\min_{\beta \in \mathbb{R}^p} \| \mathbf{Y} - \beta \|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} \mathbf{1} \left( \beta_{i+1} \neq \beta_i \right) \le k$$

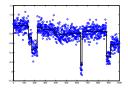
- This is an optimization problem over the (<sup>P</sup><sub>k</sub>) partitions...
   Dynamic programming finds the solution in O(p<sup>2</sup>k) in time and O(p<sup>2</sup>) in memory
- But: does not scale to  $p = 10^6 \sim 10^9$ .



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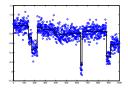
- This is an optimization problem over the  $\binom{p}{k}$  partitions...
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For a signal Y ∈ ℝ<sup>ρ</sup>, define an optimal approximation β ∈ ℝ<sup>ρ</sup> with k breakpoints as the solution of

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# Promoting sparsity with the $\ell_1$ penalty

#### The $\ell_1$ penalty (Tibshirani, 1996; Chen et al., 1998)

If  $R(\beta)$  is convex and "smooth", the solution of

$$\min_{\beta \in \mathbb{R}^p} \boldsymbol{R}(\beta) + \lambda \sum_{i=1}^p |\beta_i|$$

Geometric interpretation with p = 2

The total variation / variable fusion penalty

If  $R(\beta)$  is convex and "smooth", the solution of

$$\min_{\beta \in \mathbb{R}^p} R(\beta) + \lambda \sum_{i=1}^{p-1} |\beta_{i+1} - \beta_i|$$

is usually piecewise constant (Rudin et al., 1992; Land and Friedman, 1996).

Proof:

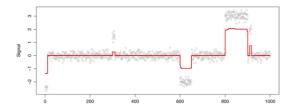
- Change of variable  $u_i = \beta_{i+1} \beta_i$ ,  $u_0 = \beta_1$
- We obtain a Lasso problem in  $u \in \mathbb{R}^{p-1}$
- u sparse means  $\beta$  piecewise constant

# TV signal approximator

$$\min_{\beta \in \mathbb{R}^p} \| Y - \beta \|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} |\beta_{i+1} - \beta_i| \le \mu$$

Adding additional constraints does not change the change-points:

•  $\sum_{i=1}^{p} |\beta_i| \le \nu$  (Tibshirani et al., 2005; Tibshirani and Wang, 2008) •  $\sum_{i=1}^{p} \beta_i^2 \le \nu$  (Mairal et al. 2010)



$$\min_{\beta \in \mathbb{R}^p} \| Y - \beta \|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} |\beta_{i+1} - \beta_i| \le \mu$$

- QP with sparse linear constraints in O(p<sup>2</sup>) -> 135 min for p = 10<sup>5</sup> (Tibshirani and Wang, 2008)
- Coordinate descent-like method O(p)? -> 3s s for p = 10<sup>5</sup> (Friedman et al., 2007)
- With the LARS in *O*(*pk*) (Harchaoui and Levy-Leduc, 2008)
- For all  $\mu$  in  $O(p \ln p)$  (Hoefling, 2009)
- For the first k change-points in  $O(p \ln k)$  (Bleakley and V., 2010)

# Solving TV signal approximator in $O(p \ln k)$

#### Theorem (V. and Bleakley, 2010; see also Hoefling, 2009)

TV signal approximator performs "greedy" dichotomic segmentation

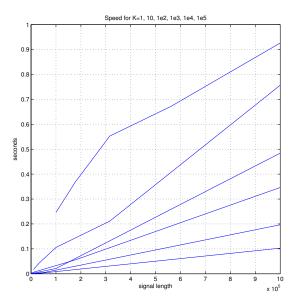
Algorithm 1 Greedy dichotomic segmentation

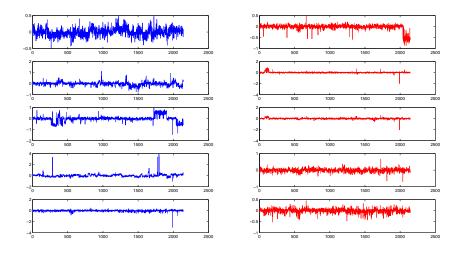
**Require:** k number of intervals,  $\gamma(I)$  gain function to split an interval I into  $I_L(I)$ ,  $I_R(I)$ 

1:  $I_0$  represents the interval [1, n]2:  $\mathcal{P} = \{I_0\}$ 3: for i = 1 to k do 4:  $I^* \leftarrow \arg \max \gamma (I^*)$ 5:  $\mathcal{P} \leftarrow \mathcal{P} \setminus \{I^*\}$ 6:  $\mathcal{P} \leftarrow \mathcal{P} \cup \{I_L (I^*), I_R (I^*)\}$ 7: end for 8: return  $\mathcal{P}$ 

#### Apparently greedy algorithm finds the global optimum!

# Speed trial : 2 s. for k = 100, $p = 10^7$



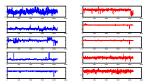


#### Aggressive (left) vs non-aggressive (right) melanoma

JP Vert (ParisTech)

#### Sparse methods in genomics

## Fused lasso for supervised classification



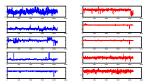
 Idea: find a linear predictor f(Y) = β<sup>T</sup> Y that best discriminates the aggressive vs non-aggressive samples, subject to the constraints that it should be sparse and piecewise constant

• Mathematically:

$$\min_{\beta \in \mathbb{R}^{p}} R(\beta) + \lambda_{1} \|\beta\|_{1} + \lambda_{2} \|\beta\|_{TV}$$

#### Computationnally: proximal methods

## Fused lasso for supervised classification

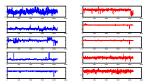


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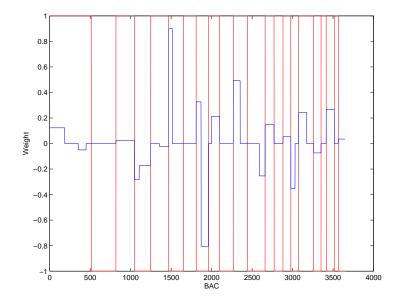
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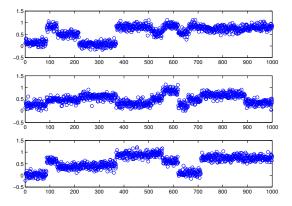
#### Computationnally: proximal methods

JP Vert (ParisTech)

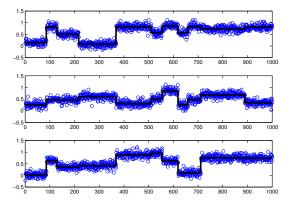
## Prognosis in melanoma (Rapaport et al., 2008)



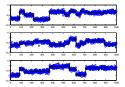
# Extension 2: finding multiple change points shared by several profiles



# Extension 2: finding multiple change points shared by several profiles



## "Optimal" segmentation by dynamic programming



• Define the "optimal" piecewise constant approximation  $\hat{U} \in \mathbb{R}^{p \times n}$  of *Y* as the solution of

$$\min_{U \in \mathbb{R}^{p \times n}} \| Y - U \|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} \mathbf{1} \left( U_{i+1, \bullet} \neq U_{i, \bullet} \right) \le k$$

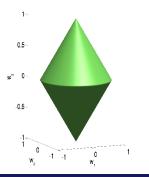
- DP finds the solution in  $O(p^2 kn)$  in time and  $O(p^2)$  in memory
- But: does not scale to  $p = 10^6 \sim 10^9$ ...

## Selecting pre-defined groups of variables

#### Group lasso (Yuan & Lin, 2006)

If groups of covariates are likely to be selected together, the  $\ell_1/\ell_2$ -norm induces sparse solutions *at the group level*:

$$\Omega_{group}(w) = \sum_{g} \|w_g\|_2$$



$$\Omega(w_1, w_2, w_3) = \|(w_1, w_2)\|_2 + \|w_3\|_2$$
$$= \sqrt{w_1^2 + w_2^2} + \sqrt{w_3^2}$$

## GFLseg (Bleakley and V., 2011)

#### Replace

$$\min_{U\in\mathbb{R}^{p\times n}} \|Y-U\|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} \mathbf{1} \left( U_{i+1,\bullet} \neq U_{i,\bullet} \right) \leq k$$

by

$$\min_{U \in \mathbb{R}^{p \times n}} \|Y - U\|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} w_i \|U_{i+1,\bullet} - U_{i,\bullet}\| \le \mu$$

GFLseg = Group Fused Lasso segmentation

#### Questions

- Practice: can we solve it efficiently?
- Theory: does it recover the correct segmentation?

## GFLseg (Bleakley and V., 2011)

#### Replace

$$\min_{U\in\mathbb{R}^{p\times n}} \parallel Y-U \parallel^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} \mathbf{1} \left( U_{i+1,\bullet} \neq U_{i,\bullet} \right) \leq k$$

by

$$\min_{U \in \mathbb{R}^{p \times n}} \|Y - U\|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} w_i \|U_{i+1,\bullet} - U_{i,\bullet}\| \le \mu$$

GFLseg = Group Fused Lasso segmentation

### Questions

- Practice: can we solve it efficiently?
- Theory: does it recover the correct segmentation?

• Make the change of variables:

$$\gamma = U_{1,\bullet},$$
  

$$\beta_{i,\bullet} = w_i \left( U_{i+1,\bullet} - U_{i,\bullet} \right) \quad \text{for } i = 1, \dots, p-1.$$

• TV approximator is then equivalent to the following group Lasso problem (Yuan and Lin, 2006):

$$\min_{\beta \in \mathbb{R}^{(p-1) \times n}} \| \bar{Y} - \bar{X}\beta \|^2 + \lambda \sum_{i=1}^{p-1} \| \beta_{i,\bullet} \|,$$

where  $\overline{Y}$  is the centered signal matrix and  $\overline{X}$  is a particular  $(p-1) \times (p-1)$  design matrix.

## TV approximator implementation

$$\min_{\beta \in \mathbb{R}^{(p-1) \times n}} \| \bar{Y} - \bar{X}\beta \|^2 + \lambda \sum_{i=1}^{p-1} \| \beta_{i,\bullet} \|,$$

#### Theorem

The TV approximator can be solved efficiently:

- "approximately" with the group LARS in *O*(*npk*) in time and *O*(*np*) in memory
- "exactly" with a block coordinate descent + active set method in O(np) in memory

Although  $\bar{X}$  is  $(p-1) \times (p-1)$ :

- For any *R* ∈ ℝ<sup>p×n</sup>, we can compute *C* = X<sup>T</sup>*R* in *O*(*np*) operations and memory
- For any two subset of indices  $A = (a_1, \dots, a_{|A|})$  and  $B = (b_1, \dots, b_{|B|})$  in [1, p-1], we can compute  $\bar{X}_{\bullet,A}^{\top} \bar{X}_{\bullet,B}$  in O(|A||B|) in time and memory
- For any  $A = (a_1, \ldots, a_{|A|})$ , set of distinct indices with  $1 \le a_1 < \ldots < a_{|A|} \le p 1$ , and for any  $|A| \times n$  matrix R, we can compute  $C = (\bar{X}_{\bullet,A}^{\top} \bar{X}_{\bullet,A})^{-1} R$  in O(|A|n) in time and memory

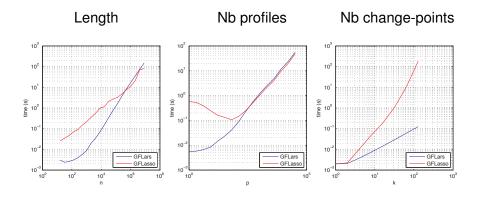
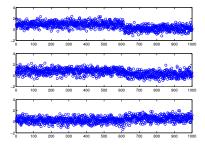


Figure 2: Speed trials for group fused LARS (top row) and Lasso (bottom row). Left column: varying n, with fixed p = 10 and k = 10; center column: varying p, with fixed n = 1000 and k = 10; right column: varying k, with fixed n = 1000 and p = 10. Figure axes are log-log. Results are averaged over 100 trials.

## Consistency

Suppose a single change-point:

- at position  $u = \alpha p$
- with increments  $(\beta_i)_{i=1,...,n}$  s.t.  $\bar{\beta}^2 = \lim_{k \to \infty} \frac{1}{n} \sum_{i=1}^n \beta_i^2$
- corrupted by i.i.d. Gaussian noise of variance  $\sigma^2$



Does the TV approximator correctly estimate the first change-point as *p* increases?

JP Vert (ParisTech)

## Consistency of the unweighted TV approximator

$$\min_{U \in \mathbb{R}^{p \times n}} \|Y - U\|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} \|U_{i+1,\bullet} - U_{i,\bullet}\| \le \mu$$

#### Theorem

The unweighted TV approximator finds the correct change-point with probability tending to 1 (resp. 0) as  $n \to +\infty$  if  $\sigma^2 < \tilde{\sigma}_{\alpha}^2$  (resp.  $\sigma^2 > \tilde{\sigma}_{\alpha}^2$ ), where

$$\tilde{\sigma}_{\alpha}^{2} = p \bar{\beta}^{2} \frac{(1-\alpha)^{2} (\alpha - \frac{1}{2p})}{\alpha - \frac{1}{2} - \frac{1}{2p}}$$

• correct estimation on  $[p\epsilon, p(1-\epsilon)]$  with  $\epsilon = \sqrt{\frac{\sigma^2}{2p\beta^2}} + o(p^{-1/2})$ .

wrong estimation near the boundaries

## Consistency of the weighted TV approximator

$$\min_{\boldsymbol{U}\in\mathbb{R}^{p\times n}} \|\boldsymbol{Y}-\boldsymbol{U}\|^2 \quad \text{such that} \quad \sum_{i=1}^{p-1} w_i \|\boldsymbol{U}_{i+1,\bullet}-\boldsymbol{U}_{i,\bullet}\| \leq \mu$$

#### Theorem

The weighted TV approximator with weights

$$\forall i \in [1, p-1], \quad w_i = \sqrt{rac{i(p-i)}{p}}$$

correctly finds the first change-point with probability tending to 1 as  $n \rightarrow +\infty$ .

- we see the benefit of increasing n
- we see the benefit of adding weights to the TV penalty

JP Vert (ParisTech)

• The first change-point  $\hat{i}$  found by TV approximator maximizes  $F_i = \|\hat{c}_{i,\bullet}\|^2$ , where

$$\hat{\boldsymbol{c}} = ar{\boldsymbol{X}}^{ op} \, ar{\boldsymbol{Y}} = ar{\boldsymbol{X}}^{ op} ar{\boldsymbol{X}} eta^* + ar{\boldsymbol{X}}^{ op} oldsymbol{W} \, .$$

•  $\hat{c}$  is Gaussian, and  $F_i$  is follows a non-central  $\chi^2$  distribution with

$$G_i = \frac{EF_i}{p} = \frac{i(p-i)}{pw_i^2}\sigma^2 + \frac{\bar{\beta}^2}{w_i^2w_u^2p^2} \times \begin{cases} i^2(p-u)^2 & \text{if } i \leq u \,, \\ u^2(p-i)^2 & \text{otherwise.} \end{cases}$$

• We then just check when  $G_u = \max_i G_i$ 

## Consistency for a single change-point

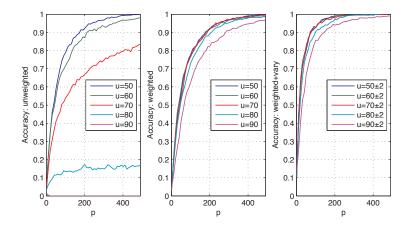


Figure 3: Single change-point accuracy for the group fused Lasso. Accuracy as a function of the number of profiles p when the change-point is placed in a variety of positions u = 50 to u = 90 (left and centre plots, resp. unweighted and weighted group fused Lasso), or:  $u = 50 \pm 2$  to  $u = 90 \pm 2$  (right plot, weighted with varying change-point location), for a signal of length 100.

## Estimation of several change-points

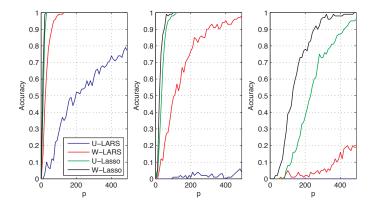
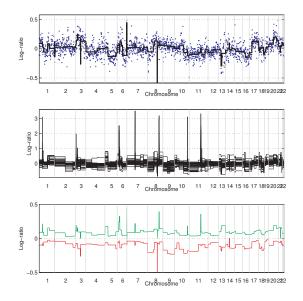


Figure 4: Multiple change-point accuracy. Accuracy as a function of the number of profiles p when change-points are placed at the nine positions  $\{10, 20, \ldots, 90\}$  and the variance  $\sigma^2$  of the centered Gaussian noise is either 0.05 (left), 0.2 (center) and 1 (right). The profile length is 100.

## Application: detection of frequent abnormalities



- Convex sparsity-inducing penalties as a way to incorporate prior knowledge
- Specific implementations for specific problems:
  - greedy dichotomic segmentation for fused lasso
  - fast group Lasso for joint segmentation
  - network flow optimization of lasso over the paths of a graph
- Often, feature selection is consistent (although we pay a price when features are very correlated), stability selection may help
- Numerous applications in bioinformatics and beyond!



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